



OPEN Construction and validation of prognostic models for young cervical cancer patients: age stratification based on restricted cubic splines

Yuan Gong^{1,2}, Feifei Gou^{1,2}, Qingfeng Qin¹, Weijie Tian¹, Wei Zhao¹ & Dan Zi¹✉

Cervical cancer (CC) ranks as the second highest cause of morbidity and mortality among young women; however, there are currently no age-specific definitions for young cervical cancer or prognostic models tailored to this demographic. Data on CC diagnosed between 2000 and 2019 were extracted from the Surveillance, Epidemiology, and End Results (SEER) database. Age stratification is based on the relationship between age and cancer-specific mortality, as demonstrated by restricted cubic spline analyses (RCS). Cox proportional hazards regression analyses were employed to identify independent prognostic factors in the young CC group. Two novel nomograms for this population were developed and validated using an external validation cohort obtained from a local hospital database, evaluated with concordance index (C-index) and calibration plots. Receiver operating characteristic (ROC) curves were utilized to compare the accuracy of the established models against the International Federation of Gynaecology and Obstetrics (FIGO) staging system (2018). A total of 27,658 patients from the SEER database were classified into three age groups (<36 years, 36-60 years, >60 years) based on RCS analyses, with 4,990, 16,922, and 5,746 patients in each group, respectively. The independent prognostic factors identified for young CC included stage, tumour size, grade, histologic type, and surgical intervention. The results of the C-index and calibration in both the training and validation sets confirmed that the two nomograms can accurately predict the occurrence and prognosis of young CC patients. The area under the curve (AUC) values indicated that these models demonstrated higher efficacy in predicting overall survival (OS) compared to the FIGO staging system (2018). These models could potentially serve as effective tools for clinicians to estimate the prognosis of young CC patients.

Cervical cancer (CC) is the fourth most frequently diagnosed cancer and the fourth leading cause of cancer death in women, with an estimated 604,000 new cases and 342,000 deaths worldwide in 2020, which represents a major global health challenge¹. Boosting rates of HPV vaccination and effective screening have reduced the incidence of CC, and the World Health Organization (WHO) is committed to eradicating cervical cancer by achieving an incidence rate of no more than 4 cases per 100 000 women-years worldwide². However, in adolescents and young women, CC remains the malignancy with the second highest morbidity and fatality rate^{3,4}; moreover, increases in the incidence or mortality of CC among young women have been reported in some regions, such as urban China, Japan, Eastern European countries and Latin America⁵⁻¹⁰.

Adolescents and young women represent a distinct demographic in cancer research. Previous studies¹¹⁻¹⁴ on young cervical cancer have encompassed a range of ages from 25 to 45 years, leading to a lack of consensus on the definition of young cervical cancer. Additionally, previous studies posited that young cervical cancer were involved in more aggressive pathogenesis, resulting in a poorer prognosis¹⁵⁻¹⁷; thus, youth was recognized as a prognostic factor for cervical cancer. However, to the best of our knowledge, there is currently a limited number of studies providing reliable data on the clinicopathological features and prognostic factors associated with young CC, and no predictive model for its prognosis has been established to date. In this study, we performed age stratification by a novel method and conducted this retrospective study to analyse the clinicopathological characteristics, treatments, and prognosis of young cervical cancer patients. We established visual prognostic

¹Department of Gynaecology, Guizhou Provincial People's Hospital, Guiyang550001, China. ²These authors contributed equally: Yuan Gong and Feifei Gou. ✉email: zidangy08@163.com

models as a supplement to International Federation of Gynaecology and Obstetrics (FIGO) staging to better predict the prognosis of young CC patients, which has not been studied before.

Methods

Database and participants

The Surveillance, Epidemiology, and End Results (SEER) database (<https://seer.cancer.gov/>) is a US population-based cancer registry database. We used SEER*Stat software (version 8.4.1) to extract information on patients diagnosed with cervical cancer between 2000 and 2019. The primary topographic site was selected using ICD codes for Endocervix (C53.0), Exocervix (C53.1), Overlapping lesions of cervix uteri (C53.8) and Cervix uteri (C53.9). The data used for external validation were obtained from the Gynaecological database of the Guizhou Provincial People's Hospital, and information on patients with cervical cancer diagnosed between 2014 and 2020 was extracted. Inclusion was based on the following: first diagnosis of cervical cancer with primary focus and diagnosis confirmed by histopathology. The exclusion criteria were as follows: (1) survival time < 1 month and (2) missing data for the chosen variables.

Variables

Risk factors used for the analysis included age, marital status at diagnosis, median household income (inflation-adjusted), months from diagnosis to treatment, grade, histologic type (squamous, adenocarcinomas, and others including uncommon subtypes such as complex epithelial neoplasms, small cell carcinoma and adenosquamous cell carcinoma), tumour size (< 2 cm, 2–4 cm, > 4 cm), stage, surgery, radiotherapy and chemotherapy. Since 2018, the pathological grading of tumours reported in the SEER database has been in accordance with a 3-category classification system, meaning that grade 4 (undifferentiated; anaplastic) was combined with grade 3 (poorly differentiated). To ease the interpretation of the results, we converted all data before 2018 into the 3-stage system. Besides, the FIGO staging system of cervical cancer was changed after 2018, mainly to include stage III lymph node metastasis. To reflect the role of lymph node metastasis in the prognosis of cervical cancer and be convenient for gynaecological clinical application, cases included were converted to 2018 FIGO stage (I, II, III, IV) in our study according to the SEER TNM staging data [AJCC 3rd (1988–2003) TNM data; Derived AJCC TNM stage, 6th edition (2004–2015); Derived SEER Combined TNM stage (2016–2017); Derived EOD 2018 TNM stage (2018+)]. In addition, according to the reported median annual household income in the United States in a previous study, \$60,000 was set as the cut-off value to group the patients¹⁸.

Outcomes of interest

Overall survival (OS) and cancer-specific survival (CSS) were set as outcomes. OS was defined as the period from the date of diagnosis until death attributed to any cause or to the end of follow-up. CSS was defined as the interval from the date of diagnosis until death as a result of cervical cancer or to the end of follow-up.

Statistical analysis

Continuous data are presented as medians and ranges, and categorical data are shown as frequencies and proportions. The Kaplan–Meier method in X-tile software (Version 3.6.1, <http://tissuearray.org>) was used to evaluate the optimal cut-off value for months from diagnosis to treatment (Supplement 1). The relationship of age and the cancer-specific mortality of CC was fit by univariate Cox regression with restricted cubic spline (RCS) analyses (knot = 5)¹⁹. The cut-off values were set for age stratification in CC patients based on the inflection points in RCS. OS and CSS of different age groups were estimated using the Kaplan–Meier method. Univariate Cox regression analysis was used to calculate the risk ratio (hazard ratio, HR) and the associated 95% confidence interval (CI) to screen potential prognostic factors affecting the OS and CSS of young CC patients from the SEER database; variables with statistically significant differences (p value < 0.05) were included in the multivariate Cox regression model to screen for independent risk factors (p value < 0.01) for prognosis in young CC patients. A nomogram model was constructed based on the results of the multivariate Cox regression analysis and assessed using the external validation cohort. The predictive ability of the model for death was assessed by the concordance index (C-index), and a higher C-index indicates a better ability to separate patients with different survival outcomes. The slopes of the calibration curves were used to compare the predicted probability with the observed probability in the study cohort. The area under the receiver operating characteristic (ROC) curve (AUC) was used to compare the difference in the prediction accuracy between the model and the FIGO staging system (2018). All analyses were performed using R software, version 4.2.2 (R Project for Statistical Computing).

Ethical statement

We obtained signed authorization and permission to access and use the data from the SEER database and followed the protocol throughout the process to protect patient privacy. Furthermore, we retrospectively collected data from medical record system of Guizhou Provincial People's Hospital, and all patients' personal information was anonymized, so the ethical review board of Guizhou Provincial People's Hospital approved the ethical requirements for this study and waived the requirement of obtaining written informed consent from patients. This study was conducted in accordance with the revised Declaration of Helsinki.

Result

Age-stratified incidence and outcomes of CC

A total of 27,658 cases of cervical cancer from the SEER database were first enrolled in our study based on the inclusion and exclusion criteria (Supplement 2). The RCS presented a nonlinear relationship between age and cervical cancer-specific mortality (p < 0.001) (Fig. 1). The plot showed a reduction trend of risk within the lower

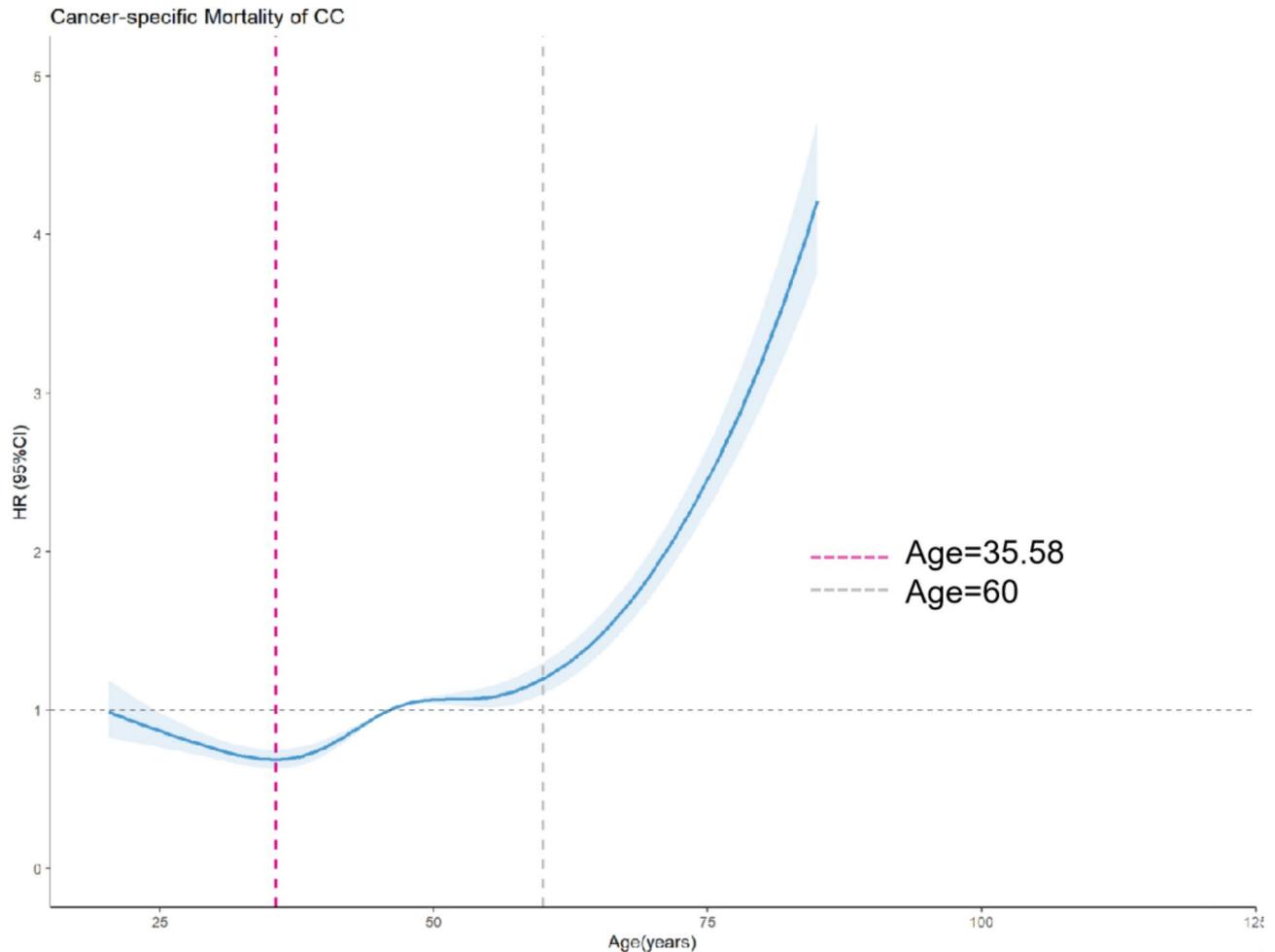


Fig. 1. The relationship of age and the cancer specific mortality of CC fit by univariate Cox regression with RCS analyses. The Cox regression indicated a nonlinear relationship between age and the risk of death ($p < 0.001$), with the lowest risk of mortality at age = 35.58 years. When the age was older than 35.58 years, HR was positively correlated with age, whereas negative correlations between age and HR were identified when age was less than 35.58 years. When age was older than 60 years, the risk of death increased substantially with increasing age. Shaded areas represent 95% CI.

age range; the lowest risk was reached at age 35.58 years old [HR = 0.687, 95% CI (0.631, 0.749)], and then risk increased thereafter with a substantial increase after 60 years old. Based on these results, 36 and 60 years were set as the cut-off values for age stratification in our study. The KM curves (Fig. 2) present the OS and CSS of cervical cancer patients in three different age groups (< 36, 36–60, > 60). Patients in the young group had the best prognosis, followed by those in the middle-aged group, and the 5-year OS and CSS in both groups were over 70%. Undoubtedly, patients in the elderly group had the worst prognosis ($P < 0.001$), with a 5-year OS rate of approximately 50%. After 5 years of cervical cancer diagnosis and treatment, the CSS KM curve of all groups tended to be flat, which also verified the regularity of disease progression and confirmed the necessity of a standardized follow-up visit.

Baseline characteristics of young group patients and differences from the other two groups

4,990 cervical cancer patients (18.0%) younger than 36 years old were assigned to young group. In terms of demographic characteristics, nearly half of the young patients were never married (44.4%), and the proportion of patients with a relatively lower median household income (<\$60,000) in this group was higher than that in the other groups (29.9% vs. 28.7% vs. 27.3%). Those patients exhibited a high compliance with treatment after diagnosis, with a higher proportion accepting treatment within 1 month (42.6% vs. 38.6% vs. 31.6%).

Regarding the clinical characteristics, in the young group, the proportion of other histologic types (nonsquamous cell neoplasms and nonadenocarcinoma) was higher than that in the other two groups (8.3% vs. 7.2% vs. 7.5%). The percentages of early-stage (FIGO stage I), grade 1 and smaller tumour sizes (< 2 cm) were higher than those of the other two groups (62.3% vs. 52.3% vs. 38.6%, 16.6% vs. 14.9% vs. 9.4%, 39.3% vs. 29.8% vs. 20.1%, respectively). The demographic and clinical characteristics of the different groups are summarized in Table 1.

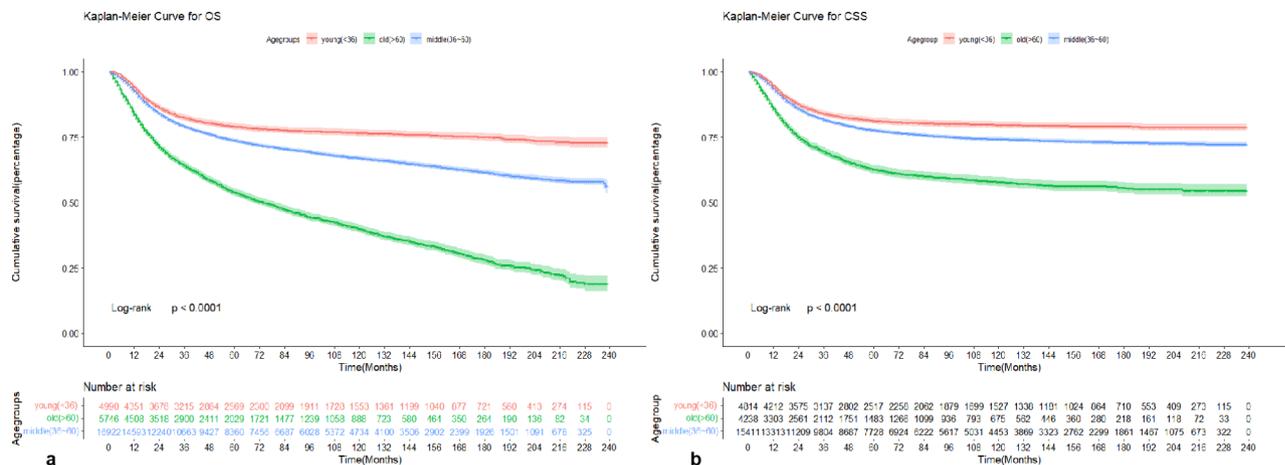


Fig. 2. Kaplan-Meier curves of OS (a) and CSS (b) for 3 groups of CC patients.

Prognostic factors for young cervical cancer patients

The young age group was used as a training cohort for developing prognostic models for young, early-onset cervical cancer. Univariate and multivariate Cox regression analyses were successively performed to identify significant risk factors in the training cohort (Table 2; Fig. 3). Pathologic type, FIGO stage, tumour size and surgery were included in the final regression models. According to the multivariate Cox regression analysis results of OS, the grade did not meet the qualification of $p < 0.01$, but as a recognized prognostic factor, it was also incorporated into the final model. In addition, chemotherapy and radiotherapy status were significant high-risk factors with $p < 0.01$ but were not included in either model because patients who received chemotherapy or radiotherapy were also worse in tumour stage or grade, which might lead to logical errors or issues of multicollinearity in the models.

Visualization and validation of the prognostic models

Nomograms for OS and CSS respectively were presented for visualization of the prognostic models (Fig. 4). Apparently, stage had the greatest effect on prognosis, followed by tumour size. Contrary to common understanding, squamous carcinoma was associated with a worse prognosis than adenocarcinoma, and both were better than other types (nonsquamous cell neoplasm and nonadenocarcinoma). The C-indexes of the OS and CSS prognosis were 0.805 (95% CI: 0.793–0.817) and 0.820 (95% CI: 0.808–0.832), respectively.

A total of 125 patients from the gynaecological database of Guizhou Provincial People's Hospital were included as the validation cohort in our study based on the inclusion and exclusion criteria. The demographic and clinical characteristics of this cohort were presented in Table S1. In the validation cohort, since all patients died from cervical cancer, we studied OS only, and the C-index was 0.865 (95% CI: 0.800–0.930).

In addition, the calibration plots of the training cohort and validation cohort showed a high fit between the nomogram-predicted survival and actual survival in terms of the probability of 3- and 5-year OS and CSS (Fig. 5).

Comparison of the nomogram with the FIGO staging system

The ROC curve was used to compare the accuracy between the established model and the FIGO staging system (2018). We plotted the ROC curves for the 3- and 5-year OS in the external validation cohort and calculated the AUC values (Fig. 6). The results showed that our model had higher efficacy in predicting OS than the FIGO staging system (2018).

Discussion

At present, the prognostic models of cervical cancer are gradually refined into special pathological subtypes, stage or disease status (such as lymph node metastasis) and special populations, and the accuracy of the models is excellent^{11,20,21}. A prognostic model for young, early-onset cervical cancer has not been reported. The question is how to define 'young' patients. Previous reports have studied patients aged 25 years or younger as a very young group because the clinicopathological characteristics and pathogenesis of this population may be different from those of adult women over 25 years old^{12,13}. There are also studies setting younger patients as under 30, 35 or 40 years with no explanation¹⁴. The results regarding the relationship between age and mortality in cervical cancer were inconsistent in studies with different samples, which may also point to a special relationship between age and prognosis. In this study, we applied a novel method to define the young group and performed age stratification for cervical cancer. RCS models have been proven to be a good tool to fit the nonlinear association between continuous variables and outcomes, and it has been confirmed that RCS, together with the Cox proportional hazards regression model, enables visualization of the shapes of dose-response associations between a continuous variable and outcome risks^{22,23}. The RCS plot presented a nonlinear relationship between age and mortality of CC patients in our study. Age of 36 and 60 were the inflection points in the plot; thus, it was

Subject	< 36	36 ~ 60	> 60	overall
	(N = 4990)	(N = 16922)	(N = 5746)	(N = 27658)
Age				
Median [Min, Max]	31.0 [14.0, 35.0]	46.0 [36.0, 60.0]	68.0 [61.0, 85.0]	47.0 [14.0, 85.0]
Marital status				
Married	2295 (46.0%)	8686 (51.3%)	2147 (37.4%)	13,128 (47.5%)
Single(never married)	2215 (44.4%)	4885 (28.9%)	910 (15.8%)	8010 (29.0%)
Other	480 (9.6%)	3351 (19.8%)	2689 (46.8%)	6520 (23.6%)
Median household income				
<\$60,000	1494 (29.9%)	4850 (28.7%)	1570 (27.3%)	7914 (28.6%)
>=\$60,000	3496 (70.1%)	12,072 (71.3%)	4176 (72.7%)	19,744 (71.4%)
Histologic subtype				
Squamous cell neoplasms	3308 (66.3%)	10,945 (64.7%)	3950 (68.7%)	18,203 (65.8%)
Adenocarcinoma	1269 (25.4%)	4762 (28.1%)	1366 (23.8%)	7397 (26.7%)
Other	413 (8.3%)	1215 (7.2%)	430 (7.5%)	2058 (7.4%)
Tumour size (cm)				
< 2	1961 (39.3%)	5049 (29.8%)	1154 (20.1%)	8164 (29.5%)
2 ~ 4	1437 (28.8%)	4959 (29.3%)	1919 (33.4%)	8315 (30.1%)
> 4	1592 (31.9%)	6914 (40.9%)	2673 (46.5%)	11,179 (40.4%)
Surgery				
No	1101 (22.1%)	5380 (31.8%)	2662 (46.3%)	9143 (33.1%)
YES	3889 (77.9%)	11,542 (68.2%)	3084 (53.7%)	18,515 (66.9%)
Stage				
I	3108 (62.3%)	8854 (52.3%)	2220 (38.6%)	14,182 (51.3%)
II	461 (9.2%)	2401 (14.2%)	1230 (21.4%)	4092 (14.8%)
III	1148 (23.0%)	3988 (23.6%)	1442 (25.1%)	6578 (23.8%)
IV	273 (5.5%)	1679 (9.9%)	854 (14.9%)	2806 (10.1%)
Months from diagnosis to treatment				
< 1	2124 (42.6%)	6531 (38.6%)	1815 (31.6%)	10,470 (37.9%)
>=1	2866 (57.4%)	10,391 (61.4%)	3931 (68.4%)	17,188 (62.1%)
Radiation				
Beam radiation	1181 (23.7%)	4920 (29.1%)	2145 (37.3%)	8246 (29.8%)
Combination	1064 (21.3%)	4616 (27.3%)	1734 (30.2%)	7414 (26.8%)
Implants	149 (3.0%)	659 (3.9%)	280 (4.9%)	1088 (3.9%)
None/Unknown	2596 (52.0%)	6727 (39.8%)	1587 (27.6%)	10,910 (39.4%)
Grade				
Grade1	826 (16.6%)	2529 (14.9%)	541 (9.4%)	3896 (14.1%)
Grade2	2134 (42.8%)	7476 (44.2%)	2376 (41.4%)	11,986 (43.3%)
Grade3	2030 (40.7%)	6917 (40.9%)	2829 (49.2%)	11,776 (42.6%)
Chemotherapy				
No/Unknown	2830 (56.7%)	8077 (47.7%)	2558 (44.5%)	13,465 (48.7%)
Yes	2160 (43.3%)	8845 (52.3%)	3188 (55.5%)	14,193 (51.3%)

Table 1. The demographic and clinical characteristics of 3 groups.

reasonable to divide patients into 3 age groups based on these two cut-off values, and patients less than 36 years old were assigned to the young group.

Previous studies suggested that younger age was associated with a worse prognosis^{16,17}. However, our findings indicate that the younger group exhibited the best prognosis, characterized by a higher proportion of early tumor stages, smaller tumor sizes, grade 1 tumors, and better medical compliance. Some other studies have reached similar conclusions^{12,14,24,25}. Regarding prognostic factors, histology type is widely acknowledged as significant in tumours. Ruiz et al. suggested that histology substantially affected OS of CC²⁴. Certain pathological subtypes of CC exhibit more aggressive behavior and are not associated with human papillomavirus (HPV) infection. HPV-negative CC was considered to be associated with a poorer prognosis^{26,27}, and the sensitivity of cytological screening for these subtypes was lower than that for cervical squamous cell carcinoma (CSCC)²⁸, potentially resulting in delays in diagnosis. In our study, the proportions of nonsquamous cell neoplasms and nonadenocarcinoma types were higher in the young group but still less than 10% (8.4%), which may have had little impact on the overall prognosis. Similarly, in other studies of cervical cancer in young patients, no significant relation was found between histological type and survival^{12,25}. In addition, the increasing incidence

Subject	Cancer-specific survival(CSS)		Overall survival(OS)	
	HR(95%CI)	p-value	HR(95%CI)	p-value
Age(years)				
< 26	Reference		Reference	
26 ~ 35	0.732 (0.588 ~ 0.910)	0.005**	0.777(0.633 ~ 0.954)	0.016*
Median household income				
<\$60,000	Reference		Reference	
>=\$60,000	0.813(0.705 ~ 0.939)	0.005**	0.811(0.712 ~ 0.924)	0.002**
Marital status				
Married	Reference		Reference	
Other	1.267(0.997 ~ 1.611)	0.053	1.292(1.039 ~ 1.605)	0.021*
Single(never married)	1.517(1.312 ~ 1.753)	< 0.001***	1.557(1.365 ~ 1.777)	< 0.001***
Months from diagnosis to treatment				
< 1	Reference		Reference	
≥ 1	1.672(1.444 ~ 1.934)	< 0.001***	1.636(1.433 ~ 1.867)	< 0.001***
Histologic type				
Adenocarcinoma	Reference		Reference	
Other	3.428(2.665 ~ 4.408)	< 0.001***	2.936(2.336 ~ 3.691)	< 0.001***
Squamous cell neoplasms	2.034(1.674 ~ 2.472)	< 0.001***	1.873(1.578 ~ 2.224)	< 0.001***
Tumour size (cm)				
< 2	Reference		Reference	
2 ~ 4	4.418(3.362 ~ 5.807)	< 0.001***	3.591(2.850 ~ 4.525)	< 0.001***
> 4	13.678(10.630 ~ 17.599)	< 0.001***	10.748(8.706 ~ 13.269)	< 0.001***
FIGO stage				
I	Reference		Reference	
II	4.053(3.209 ~ 5.119)	< 0.001***	3.689(3.005 ~ 4.529)	< 0.001***
III	5.169(4.355 ~ 6.135)	< 0.001***	4.397(3.777 ~ 5.119)	< 0.001***
IV	18.464(15.104 ~ 22.573)	< 0.001***	14.926(12.427 ~ 17.926)	< 0.001***
Grade				
1	Reference		Reference	
2	2.892(2.094 ~ 3.993)	< 0.001***	2.235(1.723 ~ 2.899)	< 0.001***
3	5.166(3.769 ~ 7.081)	< 0.001***	3.768(2.925 ~ 4.853)	< 0.001***
Surgery				
NO	Reference		Reference	
YES	0.208(0.182 ~ 0.238)	< 0.001***	0.227(0.200 ~ 0.257)	< 0.001***
Chemotherapy				
No/Unknown	Reference		Reference	
YES	6.281(5.310 ~ 7.431)	< 0.001***	5.341(4.615 ~ 6.181)	< 0.001***
Radiotherapy				
Beam radiation	Reference		Reference	
Combination	0.837(0.715 ~ 0.979)	0.026*	0.831(0.719 ~ 0.961)	0.016*
Implants	0.957(0.704 ~ 1.300)	0.778	0.944(0.709 ~ 1.258)	0.691
None/Unknown	0.154(0.127 ~ 0.186)	< 0.001***	0.187(0.158 ~ 0.220)	< 0.001***

Table 2. Univariate Cox regression analyzing the risk factors for young cervical cancer patients.

of young cervical cancer patients may be related to social factors. In this study, the proportion of single (never married) people in the young group was significantly higher than that in the other two groups, with a narrow gap in median household income compared to the middle age group. Being single and more likely to have multiple sexual partners²⁹, promotes persistent HPV infection.

The prognostic model illustrated by the visualized nomogram has been employed in clinical research across various tumours^{30–32}, including studies on cervical cancer^{33,34}. In this study, nomograms were utilized to develop a prognostic model specifically for young cervical cancer patients. The independent prognostic factors identified included stage, tumour size, grade, histologic type, and surgical intervention, consistent with the findings of previous studies^{35,36}. The relative weight of each prognostic factor was visually represented in the nomograms, demonstrating that each young cervical cancer patient with available clinical data on these prognostic factors could derive a specific score from the nomograms. This score could subsequently be utilized to predict 3- year or 5- year survival outcomes. Evaluation metrics for model accuracy included the C-index and the calibration curve, among others. We utilized the SEER database and incorporated statistically significant prognostic factors

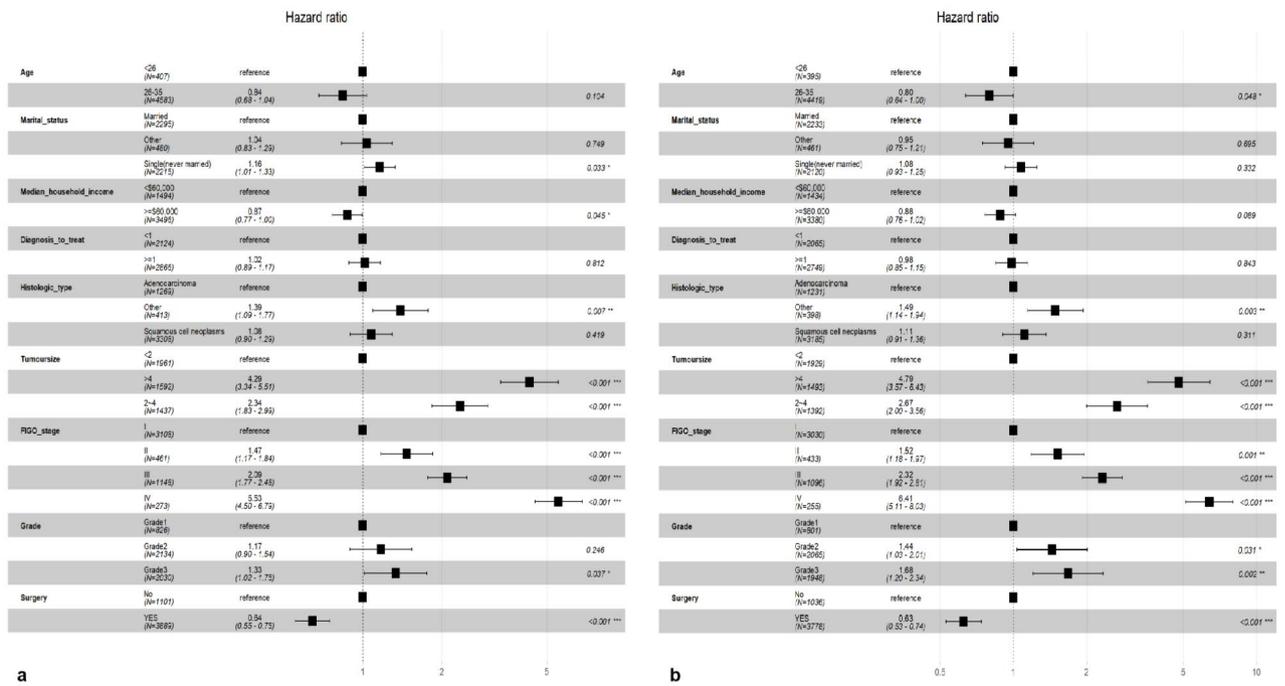


Fig. 3. Forest plots for OS (a) and CSS (b) of young CC patients based on multivariate Cox regression analysis of the training cohort.

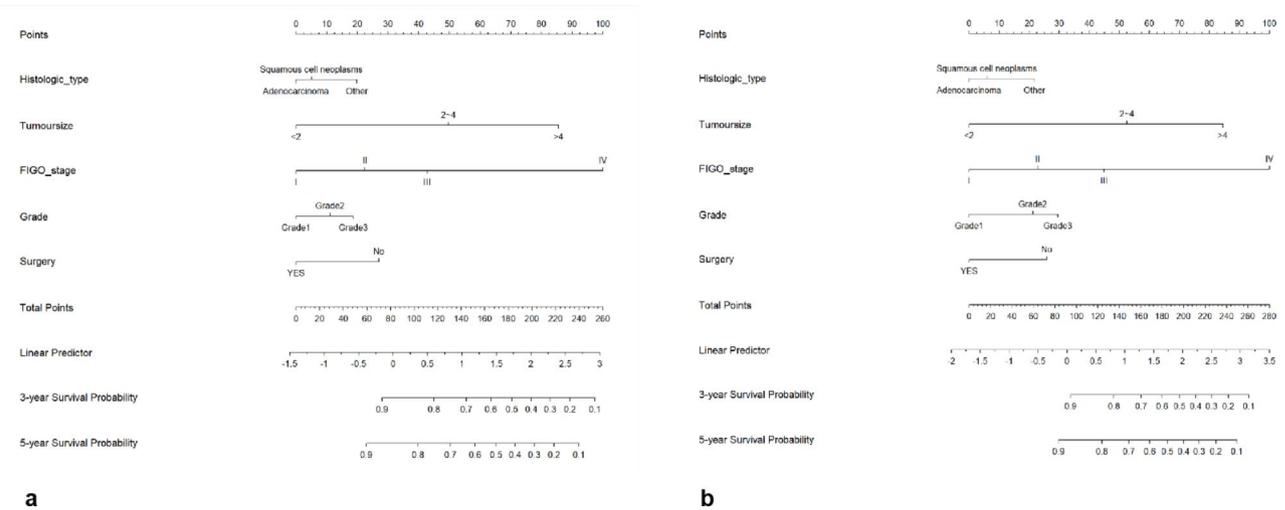


Fig. 4. Nomograms for predicting OS (a) and CSS (b) of young CC patients.

to develop a clinical prognostic model specifically for young cervical cancer patients, which has not been previously reported. The C-index for the OS model and the CSS model for young cervical cancer patients was 0.805 (95% CI: 0.793–0.817) and 0.820 (95% CI: 0.808–0.832), respectively. These values are higher than those reported for cervical cancer prognostic models in other studies^{37,38}. In the validation set, the C-index reached 0.865 (95% CI: 0.800–0.930), confirming the good predictive ability of the nomograms. The calibration plots showed a high fit between the nomogram-predicted survival and actual survival, also providing proof for the robust clinical value of the nomograms.

The FIGO staging system has been updated to the 2018 version, which has been shown to provide better predictions of disease-free survival and OS compared to earlier FIGO staging versions. In previous prognostic models, stage was typically the most significant risk factor for malignant tumours^{11,39,40}. Similarly, the FIGO stage exhibited the greatest impact on prognosis in our models. However, Grigsby et al. suggested that clinical outcomes remained heterogeneous despite having the same FIGO stage⁴¹. This observation indicated inherent limitations within the FIGO staging system. It primarily considers the extent, size, and metastasis of the tumor,

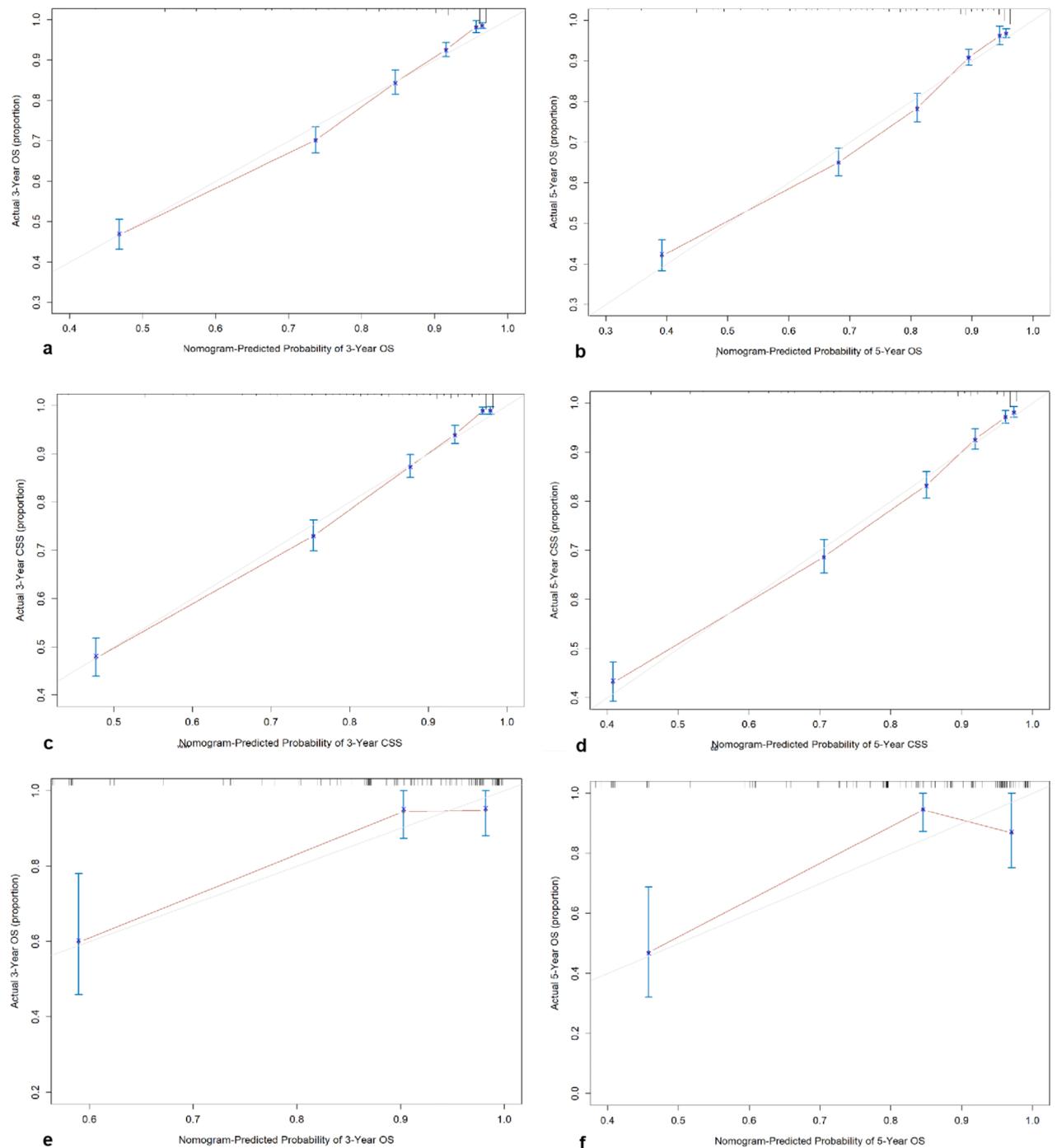


Fig. 5. Calibration plots of the models for predicting 3-, and 5-year OS and CSS of the development cohort (a–d) and 3-, and 5-year OS of external validation cohort (e, f).

while neglecting other prognostic factors such as pathological type, grade, and clinical treatment⁴². These factors were incorporated into our prognostic models, thereby enhancing the accuracy of the predictions, as evidenced by the superior AUC values for our models compared to those for the FIGO staging. Furthermore, the data for these factors are readily accessible in clinical practice.

There were some drawbacks that must be recognized in this study. Firstly, The SEER database did not contain detailed data on clinical characteristics, such as lymphovascular space invasion and HPV subtype, that are crucial for assessing the prognosis of CC. Secondly, the lack of information on lifestyle, such as age at first sexual activity, number of sexual partners and smoking history, limited the comprehensiveness of the study. Furthermore, the validation cohort was derived from a single center with a relatively small sample size ($N = 125$), which may raise

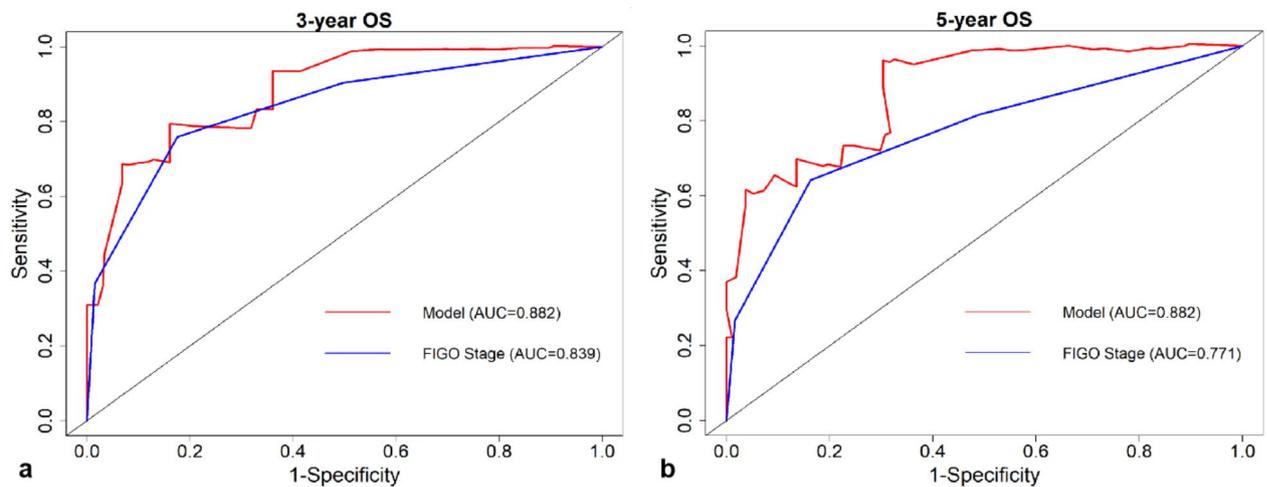


Fig. 6. ROC curves of the models and FIGO stage for predicting 3-year (a) and 5-year (b) OS in the external validation cohort.

concerns regarding potential bias. Therefore, further studies are necessary to improve accuracy and promote the model's adoption.

Conclusion

Based on the established relationship between age and mortality, we defined young CC as occurring in individuals under the age of 36. The young CC group exhibited a better prognosis compared to other age groups, with independent prognostic factors including disease stage, tumor size, grade, histologic type, and surgical intervention. Recognizing that young CC patients represent a distinct cohort with unique clinicopathological features, we developed and validated prognostic models specifically tailored for this population. These models demonstrated good calibration and high accuracy, potentially serving as a valuable supplement to the current FIGO staging system.

Data availability

The data used or analyzed during the current study are available from the corresponding author on reasonable request.

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Author contributions

D.Z. contributed to the conception of the study and review of the final manuscript; Q.Q. contributed significantly to data collection; Y.G. and W.Z. performed the data analyses and wrote the manuscript; F.G. and W.T. performed the analysis with constructive discussions. All authors approved the final manuscript.

Declarations

Competing interests

The authors declare no competing interests.

Additional information

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Correspondence and requests for materials should be addressed to D.Z.

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