



OPEN Development and validation of identification models for aortic dissection and non-ST-segment elevation acute coronary syndrome in the emergency department

Yaxin Ban^{1,4}, Hao Wang^{1,4}, Yunfei Gu¹, Haojie Chen¹, Rubing Liang¹, Yuge Jin³, Zhishuai Li¹, Xingke Li¹ & Songsen Li²✉

Aortic dissection (AD) and non-ST-segment elevation acute coronary syndrome (NSTEMI-ACS) are critical illnesses whose prompt identification within the emergency department is challenging. This study aimed to establish rapid discrimination models to differentiate between these conditions. Patients of the training set and validation set were collected from January 2020 to June 2023. All patients used their final diagnosis. Discriminant models were constructed via univariate and multivariate logistic regression analyses. Based on the results of the two models, two web calculators were developed. A total of 1314 patients were included in the study, with 997 patients (399 AD patients and 598 NSTEMI-ACS patients) and 317 patients (132 AD patients and 185 NSTEMI-ACS patients) in the training and validation sets, respectively. The semi-model consisted of six clinical characteristics (age, heart rate, pulse pressure, temperature, hypertension, diabetes), with an area under the ROC curve (AUC) of 0.792 and 0.823 in the training and validation sets. The whole-model included five clinical characteristics (age, pulse pressure, hypertension, diabetes) and two point-of-care test data (high sensitivity troponin I, D-dimer). It had a higher predictive value compared to the semi-model, with AUCs of 0.973 and 0.980 in the training and validation sets, respectively. Given the optimal cutoff point, the semi-model demonstrated a sensitivity of 0.716 and a specificity of 0.734, whereas the whole-model displayed a sensitivity of 0.930 and a specificity of 0.946. Both identification models can be used as reliable tools for rapidly identifying AD and NSTEMI-ACS.

Aortic dissection (AD) and acute coronary syndrome (ACS) are highly perilous, acute-onset, rapidly progressive conditions commonly encountered in chest pain centers^{1,2}. Clinical manifestations include chest pain, and the accompanying symptoms lack specificity. A meta-analysis revealed a misdiagnosis rate of 33.8% (562 out of 1663 cases) for AD, with ACS being the most mistaken diagnosis³. To prevent the risk of dissection rupture and bleeding, active preoperative preparation must be provided and unnecessary transportation for patients with AD must be avoided⁴. Notably, the standard treatment for ACS, which involves early dual antiplatelet therapy^{5,6}, can be fatal for AD patients and may increase the risk of interlayer rupture⁷. Therefore, promptly and accurately differentiating between the two conditions is essential, especially for non-ST-segment elevation (NSTEMI-ACS), which can be challenging to distinguish from AD given non-specific electrocardiographic manifestations. Unfortunately, no efficient identification models exist for AD and NSTEMI-ACS in emergency departments.

Therefore, based on the TRIPOD statement⁸, this study collected clinical characteristics and point-of-care test (POCT) data from patients in a chest pain center and developed semi-model (i.e., including only clinical characteristics) and whole-model (i.e., including clinical characteristics and POCT data) for patients with AD and NSTEMI-ACS. Decision curves were used to assess the clinical benefits^{9,10}, and user-friendly web calculators were developed to assist in rapid clinical diagnosis.

¹Department of Cardiology, Luoyang Central Hospital Affiliated to Zhengzhou University, Luoyang 471000, China.

²Luoyang Maternal and Child Health Hospital, No. 206 Tongqu Road, Luoyang 471000, China. ³Department of Cardiology, First People's Hospital of Xinxiang City, Xinxiang 453000, China. ⁴Yaxin Ban and Hao Wang contributed equally to this work and share first authorship. ✉email: doctorssli@sina.com

Methods

Study patients and protocol

We collected data retrospectively from patients with AD and NSTEMI-ACS at the Chest Pain Center of Luoyang Central Hospital between Jan 2020 and Jun 2022 for the training set, and prospectively collected the data from Jul 2022 to Jun 2023 for the validation set. All patients used their final diagnosis. The prediction model intends to utilize up to eight candidate predictors, with a C statistic of 0.75. The ratio of AD patients to those with NSTEMI-ACS in the chest pain center, as determined from the preliminary experiment, is 2:3. Consequently, the final calculated minimum sample size required for this study is 590 cases, and the sample size for the training set meets the established standards. The AD cohort included patients with both Stanford Type A and Type B aortic dissections, confirmed by computed tomography angiography (CTA). The diagnostic criteria for NSTEMI-ACS conform to the 2020 guidelines of the European Society of Cardiology for managing patients experiencing ACS in the absence of sustained ST-segment elevation. These criteria have been verified through coronary angiography (CAG). Exclusion criteria included: (1) Severe infection. (2) Severe liver and kidney dysfunction (Child-Pugh grade B or C, decompensated renal function, or dialysis use), malignant tumors. (3) Long-term anticoagulant drug use. (4) Refusal of CTA and CAG examinations. This study was approved by the Ethics Committee of Luoyang Central Hospital Affiliated to Zhengzhou University, and complied with the Declaration of Helsinki. All data were analyzed anonymously. Patients in the prospective cohort provided written informed consent before being included in the study. For the retrospectively collected data used in this study, patient informed consent was waived by the Ethics Committee due to the retrospective nature of the study and the use of anonymized data.

Collection of data

Demographic data, clinical features, medical history, and laboratory data—from POCT equipment—were collected from all patients during their first medical contact. Demographic information included gender and age. Clinical features included respiration, heart rate, blood pressure, and body temperature; mean arterial pressure (MAP) and pulse pressure were calculated. Medical history included hypertension, diabetes, and ischemic stroke. The diagnosis of ischemic stroke, based on patient history, involves clinical assessment and imaging confirmation (MRI). Diagnostic criteria corresponded to relevant guidelines^{11–13}. Laboratory data, including D-dimer and high sensitivity troponin I (HsTnI), were obtained through POCT equipment. They were drawn at the time of the first blood draw in the emergency department, immediately upon arrival. All data were double-entered and compared using Epidata (version 3.1).

Statistical analysis

Figure 1 demonstrates the overall design of the experiment. In the training set, both the semi-model and the whole-model were constructed, and subsequently verified in the verification set. The Shapiro–Wilk test was used to analyze the normality of continuous variables. Mean \pm standard deviation represented variables with a normal distribution. To evaluate the uniformity of variances, Levene's test was used. When variances were equivalent, the Student's t-test was applied, whereas the Welch's t-test was utilized in cases with unequal variances. For variables that deviated from a normal distribution, medians and inter quartile ranges (IQR) were employed, and group disparities were assessed through the Mann-Whitney U test. For categorical variables, the chi-square test was employed to analyze differences between groups. All hypothesis tests were two-sided and assumed a significance level of 0.05.

Construction and validation of identification models

To perform missing value imputation, multiple imputations using R's mice package were employed with the "pmm" interpolation method and a seed of 999 was set. The variable with the highest missing proportion was HsTnI, with a missing proportion of 30%. Therefore, 30 interpolations were performed¹⁴, a final complete dataset for subsequent analysis was obtained by pooling these datasets according to Rubin's rules. To evaluate the presence of multicollinearity, the car package was utilized to calculate the variance inflation factor (VIF), excluding variables showing high multicollinearity. Variables were retained for the multivariate logistic regression analysis to construct a predictive model with a P value less than 0.1. Then odds ratios (OR) and 95% confidence intervals (95%CI) were calculated. Subsequently, leveraging the findings from the multi-factor logistic regression analysis, nomograms were created to visually represent the probability of individuals experiencing AD.

To assess the calibration and discrimination of the two models, calibration and receiver operating characteristic (ROC) curves were generated in the training set. To evaluate the calibration curve, bootstrap analysis with 1000 repetitions was employed, while the ROC curve relied on the predicted scores from the nomogram. Subsequently, the area under the ROC curve (AUC) was calculated. The same methods were applied to the verification set. Decision curves were then plotted with the training and verification sets to evaluate the clinical net benefits of the models. To improve accessibility, a user-friendly web calculator was developed utilizing the Shiny package and platform. The calculator integrated the results from the two models, providing an interface for users to easily access the outcomes. All statistical analyses were performed using R software (version 4.2.3). The software packages used included mice, reportReg, rms, car, pROC, dcurves, DynNom, and Shiny.

Results

Baseline characteristics of patients

1314 participants were enrolled, with 997 and 317 individuals in the training and validation sets, respectively. Table 1 describes patient demographic information, clinical characteristics, medical background, and laboratory findings for both sets. Within the training set, the AD group exhibited higher levels of body temperature,

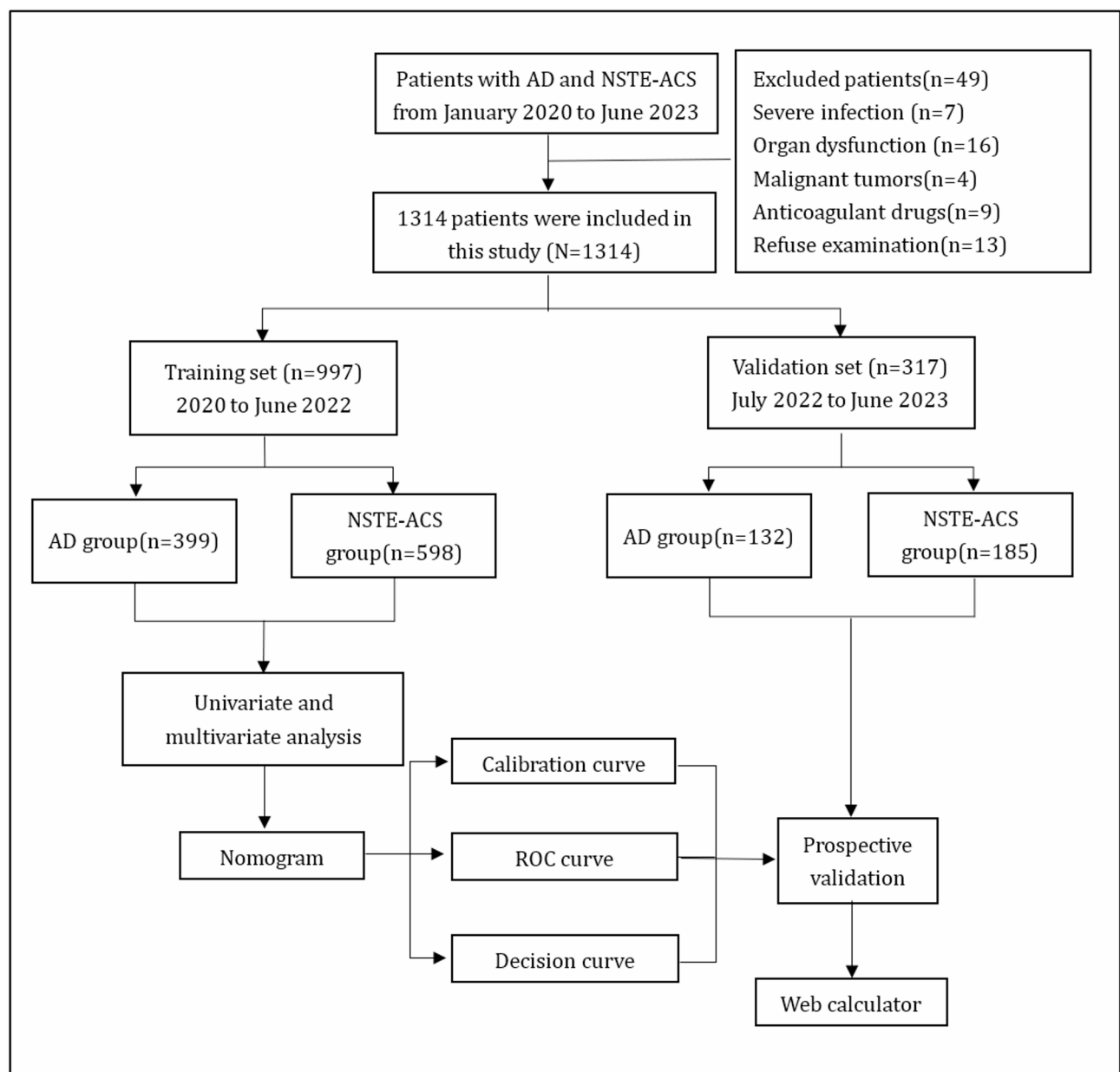


Fig. 1. Flow chart of the research. AD: Aortic dissection; NSTEMI-ACS: non-ST-segment elevation acute coronary syndrome; ROC: receiver operating characteristic.

D-dimer and greater frequencies of male patients and hypertension compared to the NSTEMI-ACS group. Conversely, the AD group displayed lower levels of HsTnI, young ages and fewer frequencies of diabetes and stroke history compared to the NSTEMI-ACS group. Heart rate, pulse pressure, systolic blood pressure, and MAP had no noticeable disparities between the two groups.

Feature selection

After performing the VIF calculation, multicollinearity was noted among systolic blood pressure, pulse pressure difference, and MAP. By excluding variables separately, all variables achieved the minimum VIF value (with the maximum VIF value being 1.46) when systolic blood pressure was excluded. Univariate logistic regression analysis was conducted. The outcome is presented in Table 2. To construct the whole-model, a multivariate logistic regression analysis was performed by incorporating variables with P-values < 0.1. After excluding laboratory data, the remaining variables were used in the regression analysis to create the semi-model. The OR and 95%CI for both models are presented in Table 3.

Characteristics	Training group(<i>n</i> = 997)			Validation group(<i>n</i> = 317)		
	AD(<i>n</i> = 399)	NSTE-ACS(<i>n</i> = 598)	<i>P</i> value	AD(<i>n</i> = 132)	NSTE-ACS(<i>n</i> = 185)	<i>P</i> value
Demographic data						
Male sex, <i>n</i> (%)	311(77.94)	407(68.06)	0.001	111(84.09)	133(71.89)	0.02
Age (years)	53(44–62)	64(55–73)	< 0.001	55(46–63)	65(55–72)	< 0.001
Clinical features						
Respiration rate (/min)	20(18–20)	19(18–20)	0.06	20(18–20)	19(19–20)	0.07
Heart rate (/min)	74(66–84)	75(66–86)	0.22	78(66–86)	78(70–84)	0.62
SBP (mmHg)	140(120–161.5)	140(126–160)	0.767	150(125–176)	136(120–154)	< 0.001
Pulse pressure (mmHg)	59(47–70)	55(48–70)	0.14	68(51–80)	55(47–66)	< 0.001
MAP (mmHg)	102(87–118.5)	103(93–117)	0.23	103(90–123)	99(89–108)	0.02
Body temperature (°C)	36.5(36.2–36.6)	36.4(36.2–36.5)	0.008	36.5(36.5–36.7)	36.5(36.3–36.6)	< 0.001
Medical history						
Hypertension, <i>n</i> (%)	297(74.44)	365(61.04)	< 0.001	109(82.58)	105(56.76)	< 0.001
Diabetes, <i>n</i> (%)	13(3.26)	165(27.59)	< 0.001	2(1.52)	55(29.73)	< 0.001
Stroke, <i>n</i> (%)	30(7.52)	96(16.05)	< 0.001	10(7.58)	30(16.22)	0.04
Laboratory data						
HsTnI (ng/mL)	0.032(0.005–0.138)	1.701(0.139–6.840)	< 0.001	0.006(0.001–0.730)	0.514(0.034–5.029)	< 0.001
D-dimer (ng/mL)	3093(1751–4745)	140(80–310)	< 0.001	3920(2274–7620)	190(160–220)	< 0.001

Table 1. Baseline characteristics of Patients. AD: Aortic dissection; NSTE-ACS: non-ST-segment elevation acute coronary syndrome; SBP: Systolic blood pressure; MAP: Mean arterial pressure; HsTnI: high sensitivity troponin I.

Variables	OR (95% CI)	<i>P</i> value
Male sex	1.659(1.238–2.222)	0.001
Age	0.940(0.930–0.951)	< 0.001
Respiration	1.041(0.979–1.107)	0.20
Heart rate	0.992(0.984–1.000)	0.06
Pulse pressure	1.006(0.999–1.013)	0.08
MAP	0.998(0.991–1.004)	0.49
Body temperature	2.186(1.315–3.633)	0.003
Hypertension	1.859(1.407–2.456)	< 0.001
Diabetes	0.088(0.049–0.158)	< 0.001
Stroke	0.425(0.276–0.654)	< 0.001
HsTnI	0.675(0.618–0.737)	< 0.001
D-dimer	1.002(1.001–1.002)	< 0.001

Table 2. Univariate logistic regression analysis of variables estimated by using the data from the training set. OR: odds ratio; MAP: mean arterial pressure; HsTnI: high sensitivity troponin I.

Construction and verification of the semi-model for AD and NSTE-ACS

A nomogram was developed using the semi-model to predict the probability of AD based on six clinical characteristics (Fig. 2). The nomogram was based on independent predictors identified in the training set. Each clinical characteristic was assigned a score ranging from 1 to 100. Individual scores were combined for the total score, which was used to estimate the probability of AD. The predicted probability aligned well with the actual occurrence probability of AD, as demonstrated by the calibration curve (Fig. 3A, B). The ROC curve (Fig. 3C, D) showed that the AUC in the training set was 0.792 (95% CI: 0.764–0.820), with a score cut-off of 143.6 proving most effective. At the optimal threshold, the sensitivity and specificity were 0.716 and 0.734, the positive predictive value (PPV) and the negative predictive value (NPV) were 0.633 and 0.801. The AUC was 0.823 (95%CI: 0.776–0.869) in the validation set. The decision curve (Fig. 3E, F) showed that the semi-model exhibited good net benefit when the threshold probability ranged from 10% to 85%. A webpage for the semi-model was developed using the Shiny package and platform (https://banyx.shinyapps.io/semi_model/).

Construction and validation of the whole-model for AD and NSTE-ACS

A nomogram was developed using the whole-model to predict the probability of AD based on six clinical characteristics. It incorporated five clinical characteristics and two laboratory test outcomes (Fig. 4); its accuracy was verified using the same method. The calibration curve (Fig. 5A, B) displayed a close alignment between

Variables	Semi-model		Whole-model	
	OR (95% CI)	P value	OR (95% CI)	P value
Male sex	0.919(0.645–1.309)	0.64	1.451(0.806–2.612)	0.214
Age	0.943(0.931–0.955)	<0.001	0.934(0.914–0.953)	<0.001
Heart rate	0.987(0.977–0.997)	0.013	0.986(0.969–1.003)	0.095
Pulse pressure	1.009(1.001–1.017)	0.04	1.017(1.004–1.031)	0.011
Body temperature	2.052(1.139–3.696)	0.02	1.258(0.466–3.400)	0.65
Hypertension	2.539(1.828–3.525)	<0.001	3.015(1.699–5.349)	<0.001
Diabetes	0.094(0.052–0.173)	<0.001	0.115(0.048–0.275)	<0.001
Stroke	0.612(0.374–1.002)	0.051	0.706(0.323–1.547)	0.385
HsTnI	-	-	0.819(0.757–0.886)	<0.001
D-dimer	-	-	1.002(1.001–1.002)	<0.001

Table 3. Multivariate logistic regression analysis of variables estimated by using the data from the training set. OR: odds ratio; HsTnI: high sensitivity troponin I.

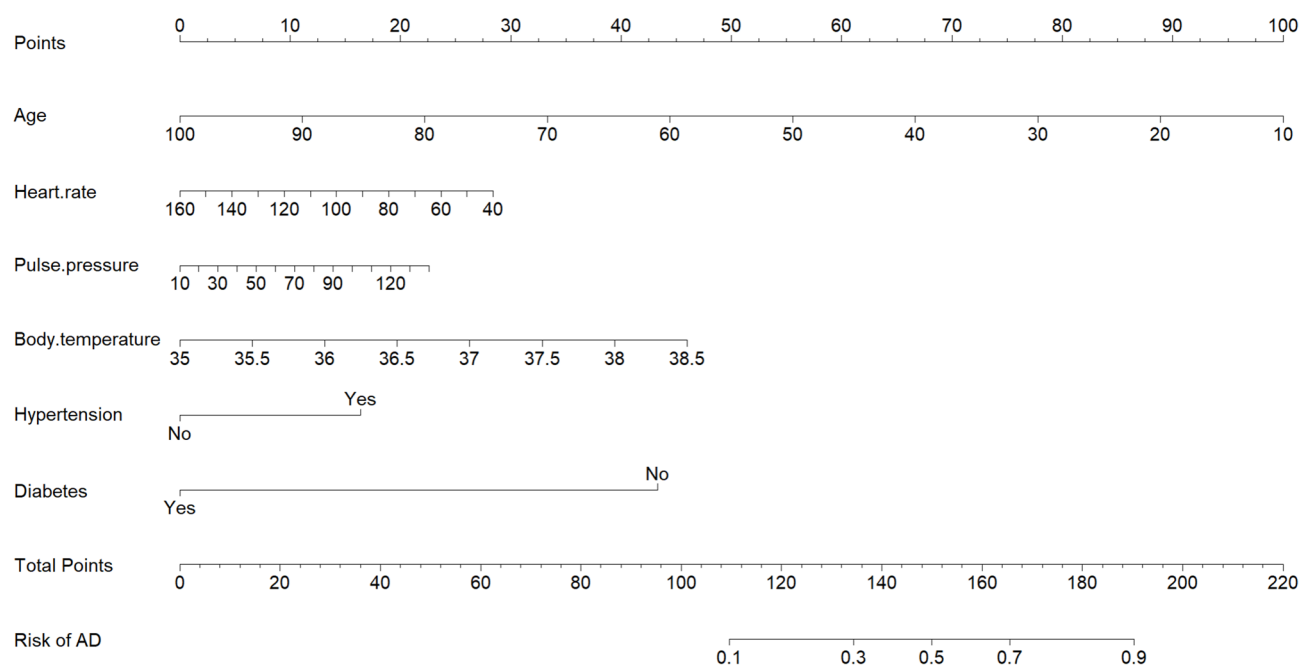


Fig. 2. Nomogram for the semi-model. AD: Aortic dissection; HsTnI: high sensitivity troponin I.

the predicted likelihood and the actual occurrence rate of AD. Moreover, the calibration curve results from the validation set surpassed those of the semi-model. The AUC was 0.973 (95%CI: 0.963–0.984) in the training set, with a score cut-off of 49.853 proving most effective. At this threshold, the sensitivity and specificity were 0.930 and 0.946, the PPV and the NPV were 0.9213 and 0.953. The AUC was 0.980 (95%CI: 0.959–1.000) in the validation set, indicating superior performance compared to the semi-model (Fig. 4C, D). The decision curves (Fig. 5E, F) illustrated that the whole-model consistently achieved better net benefit across all threshold probabilities. Finally, a web page for the whole-model (https://banyx.shinyapps.io/whole_model/) was also developed.

Discussion

This study developed and validated two identification models to differentiate between AD and NSTEMI-ACS. The semi-model was based on six clinical characteristics: age, heart rate, pulse pressure, body temperature, hypertension, and diabetes. Meanwhile, the whole-model, which incorporated laboratory test results, consisted of five clinical characteristics and two laboratory parameters: age, heart rate, pulse pressure, hypertension, diabetes, HsTnI and D-dimer. The models were prospectively verified, and ROC, calibration, and decision curves were plotted in training and validation sets. The outcomes demonstrated that both models exhibited good discrimination, calibration, and clinical net benefit. Moreover, the whole-model, showed a higher predictive value compared to the semi-model.

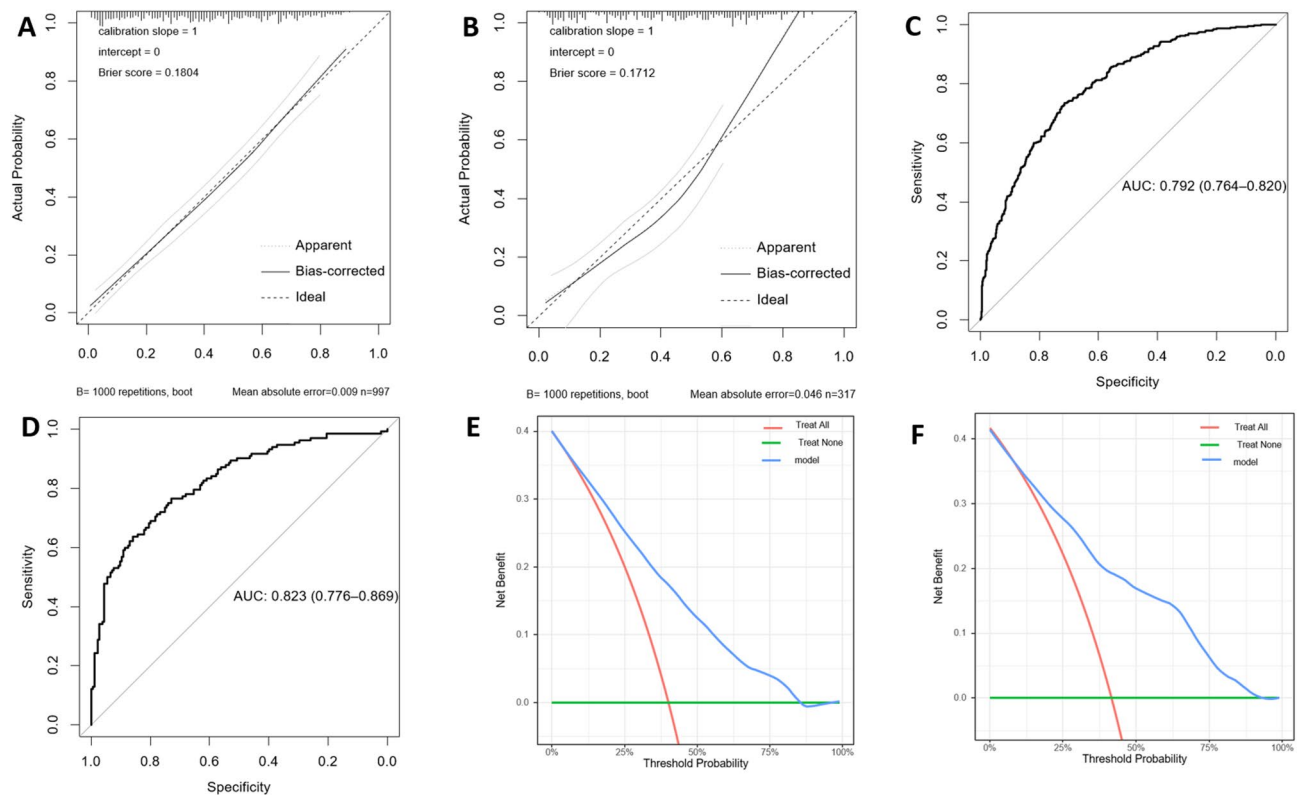


Fig. 3. Assessment of the semi-model. Calibration plots in the training set (A) and validation set (B); ROC curves in the training set (C) and validation set (D); Decision curves in the training set (E) and validation set (F); AUC: area under the receiver operating characteristic curve; ROC: receiver operating characteristic.

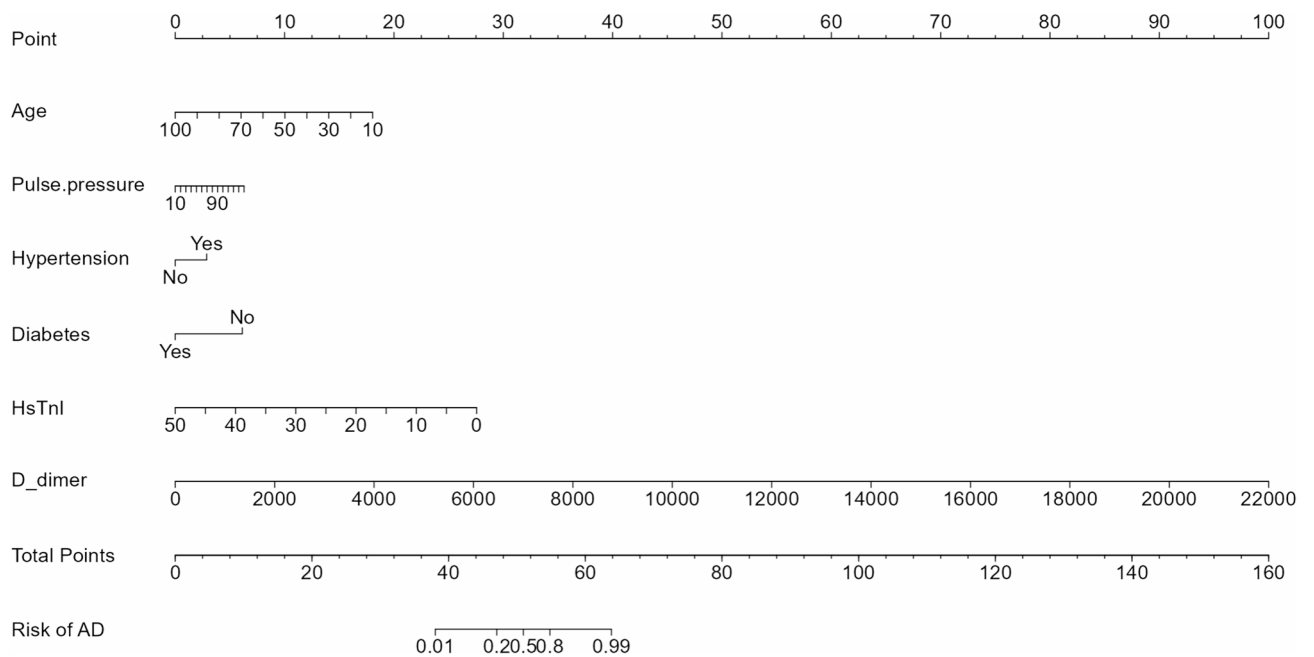


Fig. 4. Nomogram for the whole-model. AD: Aortic dissection. HsTnI: high sensitivity troponin I.

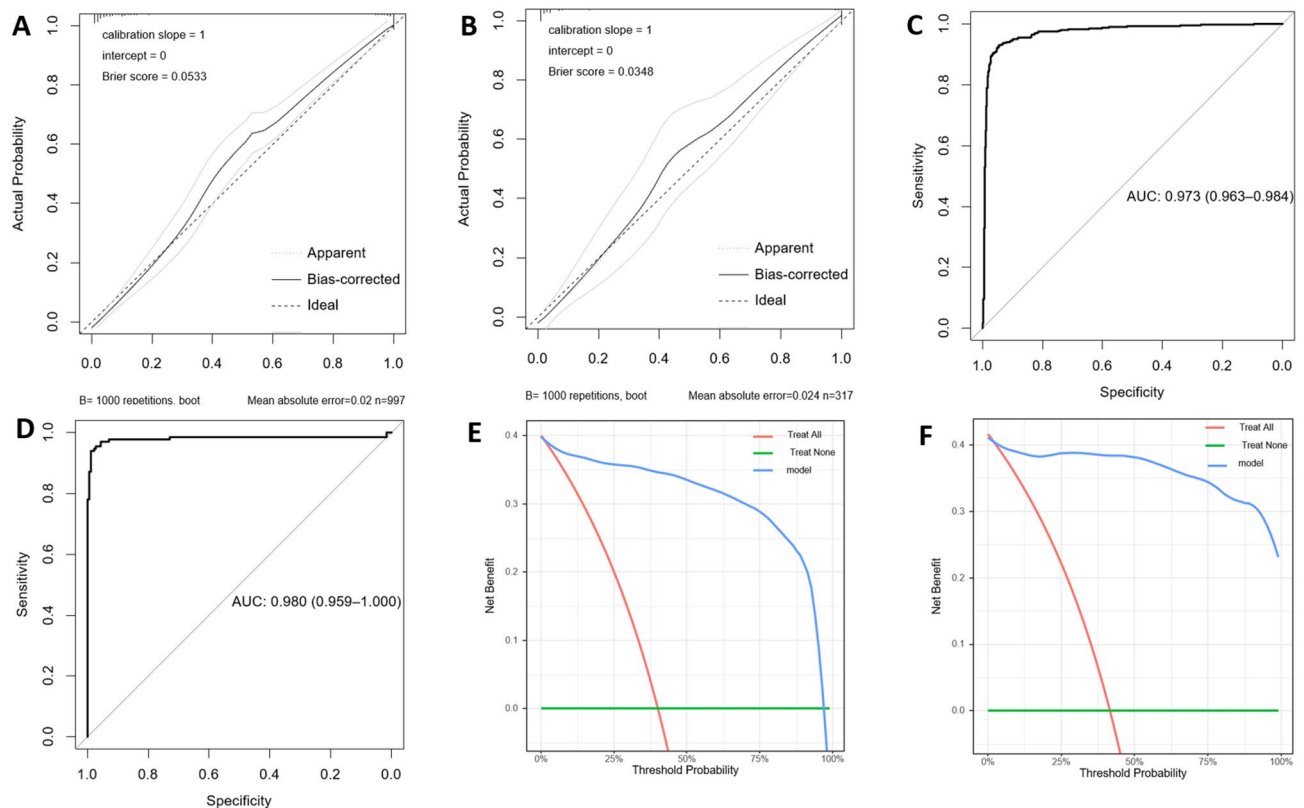


Fig. 5. Assessment of the whole-model. Calibration plots in the training set (A) and validation set (B); ROC curves in the training set (C) and validation set (D); Decision curves in the training set (E) and validation set (F); AUC: area under the curve; ROC: receiver operating characteristic; ROC: receiver operating characteristic.

Clinical prediction models, as common clinical tools, are noteworthy in the diagnoses and treatment of diseases⁸. However, few identification models exist for AD and NSTEMI-ACS. In a study¹⁵, involving 263 patients, 94 patients were diagnosed with AD, while the rest presented with NSTEMI-ACS. Training and validation sets were created following a 3:1 ratio. This study developed a model that identified four significant predictors and two clinically significant factors. While pain level and tearing pain in the model may vary based on patient tolerance, introducing subjective bias. In another investigation², a grand sum of 638 participants were divided into two groups: with chest pain less than 2 h and between 2 and 3 h. Prediction models were built for each group. However, the time window of this model was limited, making it challenging for patients to accurately determine the exact onset of the disease. This time constraint also posed difficulties for medical personnel in selecting the appropriate model. Moreover, bilirubin levels used in the model cannot be obtained through POCT. Considering these factors, further validation is required before applying these models in the emergency department.

In individuals experiencing AD, systolic hypertension aggravates the hemodynamic forces impacting the movable aortic arch and the stationary ascending and descending thoracic aortas¹⁶. Consequently, patients with AD have a higher pulse pressure. Although hypertension is also a recognized traditional risk factor for ACS, it may exert a more significant influence on the development of AD. Hoff analyzed data from the International Acute Aortic Dissection Registry and found that half of the patients with AD exhibited elevated pulse pressure¹⁷. This finding corroborates the results of the present study regarding elevated pulse pressure within the AD cohort. The study by Shimada suggested that the presence of a false lumen, progression of dissection, recurrence of dissection, or intramural thrombus may contribute to fever in patients with AD¹⁸. Inoue found that fever in AD patients could be associated with a systemic inflammatory response triggered by endogenous pro-inflammatory mediators within the body, which also affects the extrinsic coagulation pathway¹⁹. Zhang reported a significant increase in neutrophils among AD patients²⁰, further indicating the likelihood of a systemic inflammatory response in this population. These studies support the notion that body temperature elevation is prevalent among patients with AD. Of note, the variable of body temperature was not included in the whole model. The statistically significant difference in body temperature between the two groups observed in this study might be a case of significance driven by the large sample size, given that the actual difference was merely 0.1 °C. In our study, the faster heart rate observed in NSTEMI-ACS patients may be attributed to sympathetic nerve excitement. However, as heart rate was not included in the full model, its role in discriminating between the two groups requires further investigation. Lastly, advanced age and diabetes, which are traditional risk factors for ACS²¹, align with the findings of this study. These findings supported the results of the semi-model.

Biomarkers helpful diagnose AD and differentiate it from other diseases that manifest with chest pain²⁰. POCT enables the rapid detection of patient-related biomarkers at the bedside, thus benefiting patients, medical

centers, and medical systems²². Currently, D-dimer and troponin are the most frequently reported biomarkers in identifying AD and NSTEMI-ACS; both are detectable using POCT equipment²³. D-dimer is the degraded fragment of plasma fibrin after thrombus fibrinolysis²⁴. It participates in and is used as a serum marker for coagulation and fibrinolysis²⁵. In patients with AD, the interaction between blood and a non-endothelial false lumen triggers a series of coagulation and fibrinolytic reactions, increasing D-dimer levels. Numerous studies have confirmed this relationship^{26–28}, and guidelines recommend considering elevated D-dimer as a diagnostic criterion for AD^{29,30}. Nonetheless, certain investigations had also noticed heightened levels of D-dimer in individuals diagnosed with NSTEMI-ACS, albeit to a lesser degree than those with AD³¹. Additionally, patients with thrombosis also exhibited significantly elevated D-dimer levels³². These complicated the differentiation of D-dimer between AD and NSTEMI-ACS. Troponin I, a biomarker of myocardial damage, was recommended in 2000 for the diagnosis of acute myocardial infarction³³. However, several hours are required for elevated troponins to be detected in the plasma after the disease onset^{34,35}. Additionally, the level of troponin I can be normal in patients with unstable angina³⁶. The more sensitive HsTnI can be detected in plasma within one to two hours of onset³⁷ but lacks specificity and can also be elevated in some AD patients^{4,38,39}. These limitations hinder the ability of HsTnI to differentiate between the two diseases. Therefore, the current study built a whole-model, which was more effective in identifying AD and NSTEMI-ACS, based on the semi-model by incorporating the above biomarkers that can be obtained through POCT.

One advantage of the current study is that the identification model was constructed using patient data from the chest pain center, which provides a more representative sample of the population. It also had a large sample size and collected easily obtainable data. Additionally, prospective validation was conducted. The semi-model can quickly generate results, allowing for initial assessment, or guide emergency treatment before transferring patients from primary to higher-level hospitals. The whole-model can be utilized in hospitals equipped with POCT equipment to guide the management of patients with acute chest pain. Furthermore, based on these two models, we have also applied for computer software copyright (Register number: 2024SR0061185) and created a WeChat mini program called AD and non-ST-segment elevation ACS probability calculator (AppID: wx6f3d1a76133d2208), allowing clinicians and individuals to quickly calculate the probability of suspected patients suffering from AD on their cell phones or computers. This tool is valuable for identifying clinical acute chest pain patients and providing timely emergency treatment.

Limitations

Certain constraints are inherent to this research. First, the study was conducted in a specialized chest pain center within a single tertiary hospital, which may limit the generalizability of our findings to general emergency department populations. The potential for selection bias must be considered. Second, a significant limitation was the high proportion of missing data (30%) for the key predictor HsTnI in the training set. Although this was addressed using multiple imputation, this level of missingness introduces a degree of uncertainty into the model's coefficients and performance that cannot be entirely eliminated. Third, this study did not document the time between symptom onset and blood collection. A difference in this interval between the two groups could potentially lead to bias in the experimental results. Finally, this investigation relied on data from a single center. While both models demonstrate promising clinical value in prospective verification, further validation of external applicability in other centers is required.

Conclusion

In conclusion, this study developed and validated two identification models for distinguishing between AD and NSTEMI-ACS. These models can serve as effective tools for rapidly identifying the two conditions in standard chest pain centers and primary hospitals.

Data availability

Data are available from the corresponding author upon reasonable request.

Received: 7 October 2025; Accepted: 1 December 2025

Published online: 14 December 2025

References

- Bautz, B. & Schneider, J. I. High-Risk chief complaints I: chest Pain-The big three (an Update). *Emerg. Med. Clin. North. Am.* **38**, 453–498. <https://doi.org/10.1016/j.emc.2020.01.009> (2020).
- Tong, F., Wang, Y. & Sun, Z. Development and validation of nomogram models to discriminate between acute aortic syndromes and non-ST-elevation myocardial infarction during troponin-blind period. *Front. Cardiovasc. Med.* **10**, 1077712. <https://doi.org/10.3389/fcvm.2023.1077712> (2023).
- Lovatt, S. et al. Misdiagnosis of aortic dissection: A systematic review of the literature. *Am. J. Emerg. Med.* **53**, 16–22. <https://doi.org/10.1016/j.ajem.2021.11.047> (2022).
- Salmasi, M. Y. et al. Diagnosis and management of acute Type-A aortic dissection in emergency departments: results of a UK National survey. *Int. J. Cardiol.* **300**, 50–59. <https://doi.org/10.1016/j.ijcard.2019.09.074> (2020).
- Amsterdam, E. A. et al. 2014 AHA/ACC guideline for the management of patients with non-ST-elevation acute coronary syndromes: executive summary: a report of the American college of Cardiology/American heart association task force on practice guidelines. *Circulation* **130**, 2354–2394. <https://doi.org/10.1161/cir.000000000000133> (2014).
- Collet, J. P. et al. 2020 ESC guidelines for the management of acute coronary syndromes in patients presenting without persistent ST-segment elevation. *Eur. Heart J.* **42**, 1289–1367. <https://doi.org/10.1093/eurheartj/ehaa575> (2021).
- Xue, Y. et al. Aortic dissection patients mimic acute coronary syndrome with preoperative antiplatelet therapy. *J. Thorac. Dis.* **11**, 3385–3390. <https://doi.org/10.21037/jtd.2019.08.02> (2019).
- Collins, G. S., Reitsma, J. B., Altman, D. G. & Moons, K. G. Transparent reporting of a multivariable prediction model for individual prognosis or diagnosis (TRIPOD): the TRIPOD statement. *Bmj* **350**, g7594. <https://doi.org/10.1136/bmj.g7594> (2015).

9. Rousson, V. & Zumbun, T. Decision curve analysis revisited: overall net benefit, relationships to ROC curve analysis, and application to case-control studies. *BMC Med. Inf. Decis. Mak.* **11**, 45. <https://doi.org/10.1186/1472-6947-11-45> (2011).
10. Vickers, A. J. & Elkin, E. B. Decision curve analysis: a novel method for evaluating prediction models. *Med. Decis. Mak.* **26**, 565–574. <https://doi.org/10.1177/0272989x06295361> (2006).
11. Unger, T. et al. 2020 international society of hypertension global hypertension practice guidelines. *Hypertension* **75**, 1334–1357. <https://doi.org/10.1161/hypertensionaha.120.15026> (2020).
12. Cosentino, F. et al. 2019 ESC guidelines on diabetes, pre-diabetes, and cardiovascular diseases developed in collaboration with the EASD. *Eur. Heart J.* **41**, 255–323. <https://doi.org/10.1093/eurheartj/ehz486> (2020).
13. Kleindorfer, D. O. et al. 2021 guideline for the prevention of stroke in patients with stroke and transient ischemic attack: A guideline from the American heart association/American stroke association. *Stroke* **52**, e364–e467. <https://doi.org/10.1161/str.0000000000000375> (2021).
14. White, I. R., Royston, P. & Wood, A. M. Multiple imputation using chained equations: issues and guidance for practice. *Stat. Med.* **30**, 377–399. <https://doi.org/10.1002/sim.4067> (2011).
15. Zhang, B., Wang, Y., Guo, J., Zhang, G. & Yang, B. Nomogram to differentiate between aortic dissection and non-ST segment elevation acute coronary syndrome: a retrospective cohort study. *Cardiovasc. Diagn. Ther.* **11**, 457–466. <https://doi.org/10.21037/cdt-20-935> (2021).
16. Thrumurthy, S. G., Karthikesalingam, A., Patterson, B. O., Holt, P. J. & Thompson, M. M. The diagnosis and management of aortic dissection. *Bmj* **344**, d8290. <https://doi.org/10.1136/bmj.d8290> (2011).
17. Hoff, E. et al. Pulse pressure and type A acute aortic dissection in-hospital outcomes (from the international registry of acute aortic Dissection). *Am. J. Cardiol.* **113**, 1255–1259. <https://doi.org/10.1016/j.amjcard.2013.12.037> (2014).
18. Shimada, S. et al. Fever associated with acute aortic dissection. *Circ. J.* **71**, 766–771. <https://doi.org/10.1253/circj.71.766> (2007).
19. Inoue Arita, Y. et al. A fever in acute aortic dissection is caused by endogenous mediators that influence the extrinsic coagulation pathway and do not elevate procalcitonin. *Intern. Med.* **55**, 1845–1852. <https://doi.org/10.2169/internalmedicine.55.5924> (2016).
20. Zhang, H. et al. The potential value of the neutrophil to lymphocyte ratio for early differential diagnosis and prognosis assessment in patients with aortic dissection. *Clin. Biochem.* **97**, 41–47. <https://doi.org/10.1016/j.clinbiochem.2021.08.002> (2021).
21. Rosengren, A. et al. Cardiovascular risk factors and clinical presentation in acute coronary syndromes. *Heart* **91**, 1141–1147. <https://doi.org/10.1136/hrt.2004.051508> (2005).
22. Rooney, K. D. & Schilling, U. M. Point-of-care testing in the overcrowded emergency department—can it make a difference? *Crit. Care* **18**, 692. <https://doi.org/10.1186/s13054-014-0692-9> (2014).
23. Tripoliti, E. E. et al. Point-of-Care testing devices for heart failure analyzing blood and saliva samples. *IEEE Rev. Biomed. Eng.* **13**, 17–31. <https://doi.org/10.1109/rbme.2019.2905730> (2020).
24. Soomro, A. Y., Guerchicoff, A., Nichols, D. J., Suleman, J. & Dangas, G. D. The current role and future prospects of D-dimer biomarker. *Eur. Heart J. Cardiovasc. Pharmacother.* **2**, 175–184. <https://doi.org/10.1093/ehjcvp/pvv039> (2016).
25. Wang, D. et al. The diagnostic and prognostic value of D-dimer in different types of aortic dissection. *J. Cardiothorac. Surg.* **17**, 194. <https://doi.org/10.1186/s13019-022-01940-5> (2022).
26. Suzuki, T. et al. Diagnosis of acute aortic dissection by D-dimer: the international registry of acute aortic dissection substudy on biomarkers (IRAD-Bio) experience. *Circulation* **119**, 2702–2707. <https://doi.org/10.1161/circulationaha.108.833004> (2009).
27. Shao, N. et al. The role of D-dimers in the diagnosis of acute aortic dissection. *Mol. Biol. Rep.* **41**, 6397–6403. <https://doi.org/10.1007/s11033-014-3520-z> (2014).
28. Watanabe, H. et al. Diagnostic test accuracy of D-dimer for acute aortic syndrome: systematic review and meta-analysis of 22 studies with 5000 subjects. *Sci. Rep.* **6**, 26893. <https://doi.org/10.1038/srep26893> (2016).
29. Erbel, R. et al. 2014 ESC guidelines on the diagnosis and treatment of aortic diseases: document covering acute and chronic aortic diseases of the thoracic and abdominal aorta of the adult. The task force for the diagnosis and treatment of aortic diseases of the European society of cardiology (ESC). *Eur. Heart J.* **35**, 2873–2926. <https://doi.org/10.1093/eurheartj/ehu281> (2014).
30. Isselbacher, E. M. et al. 2022 ACC/AHA guideline for the diagnosis and management of aortic disease: A report of the American heart Association/American college of cardiology joint committee on clinical practice guidelines. *Circulation* **146**, e334–e482. <https://doi.org/10.1161/cir.0000000000001106> (2022).
31. Siennicka, A. et al. Gender differences in hemostatic and inflammatory factors in patients with acute coronary syndromes: a pilot study. *J. Physiol. Pharmacol.* **69**, 91–98. <https://doi.org/10.26402/jpp.2018.1.10> (2018).
32. Ohlmann, P. et al. Diagnostic and prognostic value of Circulating D-Dimers in patients with acute aortic dissection. *Crit. Care Med.* **34**, 1358–1364. <https://doi.org/10.1097/01.Ccm.0000216686.72457.Ec> (2006).
33. Alpert, J. S., Thygesen, K., Antman, E. & Bassand, J. P. Myocardial infarction redefined—a consensus document of the joint European society of cardiology/American college of cardiology committee for the redefinition of myocardial infarction. *J. Am. Coll. Cardiol.* **36**, 959–969. [https://doi.org/10.1016/s0735-1097\(00\)00804-4](https://doi.org/10.1016/s0735-1097(00)00804-4) (2000).
34. Body, R. et al. Rapid exclusion of acute myocardial infarction in patients with undetectable troponin using a high-sensitivity assay. *J. Am. Coll. Cardiol.* **58**, 1332–1339. <https://doi.org/10.1016/j.jacc.2011.06.026> (2011).
35. Keller, T. et al. Sensitive troponin I assay in early diagnosis of acute myocardial infarction. *N Engl. J. Med.* **361**, 868–877. <https://doi.org/10.1056/NEJMoa0903515> (2009).
36. Ardissino, D. et al. Thrombin activity and early outcome in unstable angina pectoris. *Circulation* **93**, 1634–1639. <https://doi.org/10.1161/01.cir.93.9.1634> (1996).
37. Holzmann, M. J. Clinical implications of high-sensitivity cardiac troponins. *J. Intern. Med.* **284**, 50–60. <https://doi.org/10.1111/joi.12779> (2018).
38. Yang, S., Manjunath, L., Willemink, M. J. & Nieman, K. The role of coronary CT angiography for acute chest pain in the era of high-sensitivity troponins. *J. Cardiovasc. Comput. Tomogr.* **13**, 267–273. <https://doi.org/10.1016/j.jcct.2019.05.007> (2019).
39. Chaulin, A. M. Elevation mechanisms and diagnostic consideration of cardiac troponins under conditions not associated with myocardial Infarction. Part 1. *Life (Basel)* **11** <https://doi.org/10.3390/life11090914> (2021).

Acknowledgements

Our appreciation goes out to both the individuals who participated in this study and the researchers who dedicated their time.

Author contributions

Writing and revision of the manuscript: YB, HW, YG, HC, RL, YJ, ZL, XL, SL. Data collection: YB, HC, RL, YJ, ZL, XL. Data analysis: YB, HW, YG, SL.

Funding

This work was supported by the Science and Technology Plan Project of Luoyang City (2022026Y) and Henan Provincial Natural Science Foundation (232300420247).

Declarations

Competing interests

The authors declare no competing interests.

Ethics approval and consent to participate

This study was reviewed and approved by the Ethics Committee of Luoyang Central Hospital Affiliated to Zhengzhou University (approval number: LWLL-2023-11-07-01), and complied with the Declaration of Helsinki. Patients in the prospective cohort provided written informed consent before being included in the study. For the retrospectively collected data used in this study, patient informed consent was waived by the Ethics Committee due to the retrospective nature of the study and the use of anonymized data.

Additional information

Supplementary Information The online version contains supplementary material available at <https://doi.org/10.1038/s41598-025-31275-9>.

Correspondence and requests for materials should be addressed to S.L.

Reprints and permissions information is available at www.nature.com/reprints.

Publisher's note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit <http://creativecommons.org/licenses/by/4.0/>.

© The Author(s) 2025